

INCAE PhD SUMMER ACADEMY DYNAMIC GAMES IN EMPIRICAL IO

Lecture 2: Single-agent dynamic discrete choice: Estimation

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LECTURE 2: Single-agent dynamic discrete choice: Estimation

1. Data
2. Maximum Likelihood Estimation (MLE):
Nested Fixed Point (NFXP) Algorithm
3. Two-Step Hotz-Miller Methods and Finite Dependence

1. Data

DATA

- The researcher has panel data of N individuals (e.g., firms) over T periods of time.
- For each individual i and time t , the researcher observes action y_{it} and vector of state variables \mathbf{x}_{it}

$$Data = \{ y_{it}, \mathbf{x}_{it} : i = 1, 2, \dots, N ; t = 1, 2, \dots, T \}$$

- In micro-econometric applications of single-agent models, we typically have that N is large (e.g., hundreds or thousands) and T is small, i.e., short panel.

EXAMPLE: MARKET ENTRY-EXIT

- We have an industry, e.g., supermarket industry.
- Firms operate in local markets, e.g., cities, neighborhoods. That is, consumer demand, output prices, and input prices are determined at the local market level.
- In this context, an "individual" i is a combination of a "firm + local market": e.g., Walmart's entry-exit decision in North-East Toronto.
- The dataset consists of $\{y_{it}, \mathbf{x}_{it} : i = 1, 2, \dots, N; t = 1, 2, \dots, T\}$:
 - y_{it} = indicator that firm-market i is active at period t .
 - \mathbf{x}_{it} = vector of local market characteristics affecting the profit of an active firm, e.g., consumer population, average income, input prices.

2. Maximum Likelihood Estimation: NFXP ALgorithm

ESTIMATION: GENERAL IDEAS

- Given a panel dataset $\{y_{it}, \mathbf{x}_{it}\}$, we are interested in estimating the unknown parameters in the primitives $\{\pi(.), f_x(.), \delta\}$.
- Let θ be the vector of structural parameters. We distinguish three components in this vector:

$$\theta = \{ \theta_{\pi}, \theta_f, \delta \}$$

where:

θ_{π} = parameters in utility function π

θ_f = parameters in transition probability of observable state var.

ESTIMATION OF θ_f

- The parameters θ_f can be estimated separately from parameters θ_π .
- More specifically, the estimation of parameters θ_f is quite standard as it does not require solving the DP problem.
- **Example.** Market size follows an AR(1) process:

$$s_{it} = \theta_{f,0} + \theta_{f,1} s_{i,t-1} + e_{it}.$$

- We can estimate $\theta_{f,0}$ and $\theta_{f,1}$ by OLS in this AR(1) regression eq.
- More generally, given the parametric transition probability function $f_x(\mathbf{x}_{i,t+1}|y_{it}, \mathbf{x}_{it}; \theta_f)$, we can estimate θ_f by Maximum Likelihood:

$$\hat{\theta}_f = \operatorname{argmax}_{\theta_f} \ell_f(\theta_f) = \sum_{i=1}^N \sum_{t=1}^{T-1} \log f_x(\mathbf{x}_{i,t+1}|y_{it}, \mathbf{x}_{it}; \theta_f)$$

MAXIMUM LIKELIHOOD ESTIMATION OF θ_π

- We can estimate θ_π by maximizing the likelihood of (y_{i1}, \dots, y_{iT}) conditional on $(\mathbf{x}_{i1}, \dots, \mathbf{x}_{iT})$
- This log-likelihood function is:

$$\begin{aligned}
 \ell(\theta_\pi) &= \sum_{i=1}^N \log \Pr(y_{i1}, \dots, y_{iT} \mid \mathbf{x}_{i1}, \dots, \mathbf{x}_{iT}; \theta_\pi) \\
 &= \sum_{i=1}^N \log \Pr(\alpha(\mathbf{x}_{i1}, \varepsilon_{i1}; \theta_\pi) = y_{i1}, \dots, \alpha(\mathbf{x}_{iT}, \varepsilon_{iT}; \theta_\pi) = y_{iT}) \\
 &= \sum_{i=1}^N \sum_{t=1}^T \log P(y_{it} \mid \mathbf{x}_{it}; \theta_\pi)
 \end{aligned}$$

where $P(y_{it} \mid \mathbf{x}_{it}; \theta_\pi)$ is the CCP function.

- To evaluate $\ell(\theta_\pi)$ we need to solve the DP problem for this value θ_π .

NESTED FIXED POINT ALGORITHM (NFXP)

- The NFXP algorithm is a gradient iterative search method to obtain the MLE of the structural parameters.
- This algorithm nests a BHHH method (outer algorithm), that searches for a root of the likelihood equations, with a value function iteration method (inner algorithm) that solves the DP problem for each trial value of the structural parameters.
- The algorithm is initialized with an arbitrary vector of structural parameters, say θ_π^0 . At BHHH iteration $n \geq 0$:

$$\hat{\theta}_\pi^{n+1} = \hat{\theta}_\pi^n + \left(\sum_{i,t} \nabla \log P_{it}(\hat{\theta}_\pi^n) \nabla \log P_{it}(\hat{\theta}_\pi^n)' \right)^{-1} \left(\sum_{i,t} \nabla \log P_{it}(\hat{\theta}_\pi^n) \right)$$

and $\nabla \log P_{it}(\theta_\pi)$ is the gradient of log-CCP with respect to θ_π .

Nested fixed point (NFXP) algorithm (2/4)

- To illustrate this algorithm in more detail, consider a logit model where the utility function is linear in parameters:

$$\pi(y_{it}, \mathbf{x}_{it}, \boldsymbol{\theta}_\pi) = \mathbf{z}(y_{it}, \mathbf{x}_{it})' \boldsymbol{\theta}_\pi$$

where $\mathbf{z}(y_{it}, \mathbf{x}_{it})$ is a vector of known functions.

- For this model, the gradient of log-CCP is:

$$\nabla \log P_{it}(\boldsymbol{\theta}_\pi) = \mathbf{z}(y_{it}, \mathbf{x}_{it}) - \sum_{j=0}^J P(j|\mathbf{x}_{it}, \boldsymbol{\theta}_\pi) \mathbf{z}(j, \mathbf{x}_{it})$$

- The CCP function is:

$$P(y|\mathbf{x}, \boldsymbol{\theta}_\pi) = \frac{\exp \{ \mathbf{z}(y, \mathbf{x}) \boldsymbol{\theta}_\pi + \delta \mathbf{F}_x(y, \mathbf{x})' \mathbf{V}^\sigma(\boldsymbol{\theta}_\pi) \}}{\sum_{j=0}^J \exp \{ \mathbf{z}(j, \mathbf{x}) \boldsymbol{\theta}_\pi + \delta \mathbf{F}_x(j, \mathbf{x})' \mathbf{V}^\sigma(\boldsymbol{\theta}_\pi) \}}$$

Nested fixed point (NFXP) algorithm (3/4)

- The vector of values $\mathbf{V}^\sigma(\boldsymbol{\theta}_\pi)$ can be obtained as the unique fixed point of the following Integrated Bellman equation in vector form:

$$\mathbf{V}^\sigma = \log \left(\sum_{j=0}^J \exp \{ \mathbf{z}(j) \boldsymbol{\theta}_\pi + \delta \mathbf{F}_x(j) \mathbf{V}^\sigma \} \right)$$

with $\mathbf{z}(j)$ and $\mathbf{F}_x(j)$ are the matrices $\{ \mathbf{z}(j, x) : x \in X \}$ and $\{ \mathbf{F}_x(j, x) : x \in X \}$, respectively.

Nested fixed point (NFXP) algorithm (4/4)

- The NFXP algorithm works as follows.

(I) [Inner Algorithm] Given $\hat{\theta}_{\pi}^n$, we obtain the vector $\mathbf{V}^{\sigma}(\hat{\theta}_{\pi}^n)$ by successive iterations in the Integrated Bellman equation.

(II) Given $\hat{\theta}_{\pi}^n$ and $\mathbf{V}^{\sigma}(\hat{\theta}_{\pi}^n)$, we construct the CCPs $P(y_{it}|\mathbf{x}_{it}, \hat{\theta}_{\pi}^n)$ and the gradients of these CCPs using the expression above.

(III) [Outer iteration] We apply a BHHH iteration to obtain a new $\hat{\theta}_{\pi}^{n+1}$.

* We proceed in this way until the distance between $\hat{\theta}_{\pi}^{n+1}$ and $\hat{\theta}_{\pi}^n$ is smaller than a pre-specified convergence constant.

3. Two-Step Hotz-Miller Methods and Finite Dependence

MAIN IDEAS

- The cost of solving some DP problems (the "curse of dimensionality"), limits the range of applications where the NFXP can be applied.
 - Hotz and Miller (REStud, 1993) observed that, under standard assumptions in this model, it is not necessary to solve the DP problem, even once, to estimate the structural parameters.
 - Hotz-Miller approach is based on **two main ideas**:
1. **CCPs can be estimated nonparametrically** in a first-step, and these estimates can be used to construct agent's present discounted values without solving the DP problem.
 2. A large class of models have a **Finite Dependence** property. This property implies moment conditions that involve CCPs and utilities at only a small number of time periods (sometimes as small as 2).

MAIN IDEAS (2/2)

- Here I present the Finite Dependence version of Hotz-Miller approach, based on Arcidiacono & Miller (2011).
- I present this approach in the context of a dynamic multinomial logit model.

MODEL

- **Integrated Bellman equation:**

$$V^{\sigma}(\mathbf{x}_t) = \log \left[\sum_{y=0}^J \exp \{ v(y, \mathbf{x}_t) \} \right]$$

- where $v(y, \mathbf{x}_t)$ are the **Conditional Choice value functions**:

$$v(y, \mathbf{x}_t) \equiv \pi(y, \mathbf{x}_t) + \delta \sum_{\mathbf{x}_{t+1}} V^{\sigma}(\mathbf{x}_{t+1}) f_x(\mathbf{x}_{t+1} | y, \mathbf{x}_t)$$

- The **Conditional Choice Probabilities (CCPs)** are:

$$P(y | \mathbf{x}_t) = \frac{\exp \{ v(y, \mathbf{x}_t) \}}{\sum_{j=0}^J \exp \{ v(j, \mathbf{x}_t) \}}$$

FINITE DEPENDENCE PROPERTY

- **Main idea:** Under some conditions, optimal behavior implies that there is a known function that relates CCPs and utility function at periods t and $t + 1$ [more generally, at $t, t + 1, \dots, t + s$ where s is finite].

$$\mathbb{E}_t [G(\pi(y_t, \mathbf{x}_t; \theta_\pi), P(y_t | \mathbf{x}_t), \pi(y_{t+1}, \mathbf{x}_{t+1}; \theta_\pi), P(y_{t+1} | \mathbf{x}_{t+1}))] = 0$$

where $G(\cdot)$ is known.

- This equation has the same flavor as an Euler equation.
- Suppose that we can estimate the CCPs $P(y_t | \mathbf{x}_t)$ directly from the data, as reduced-form probabilities, without solving the model.
- Then, we can estimate the structural parameters in θ_π by GMM without having to solve the model even once, and without having to compute any present value.

FINITE DEPENDENCE REPRESENTATION

(1/4)

- Given the structure of the Logit CCPs, we have that:

$$\log P(0 \mid \mathbf{x}_t) = -\log \left[1 + \sum_{j=1}^J \exp \{ v(j, \mathbf{x}_t) - v(0, \mathbf{x}_t) \} \right]$$

- This implies the following expression for the integrated value function:

$$\begin{aligned} V^\sigma(\mathbf{x}_t) &= \log \left[\sum_{j=0}^J \exp \{ v(j, \mathbf{x}_t) \} \right] \\ &= v(0, \mathbf{x}_t) + \log \left[1 + \sum_{j=1}^J \exp \{ v(j, \mathbf{x}_t) - v(0, \mathbf{x}_t) \} \right] \\ &= v(0, \mathbf{x}_t) - \ln P(0, \mathbf{x}_t) \end{aligned}$$

FINITE DEPENDENCE REPRESENTATION (2/4)

- Second, for any two choice alternatives, say j and k , we have that:

$$\log P(j \mid \mathbf{x}_t) - \log P(k \mid \mathbf{x}_t) = v(j, \mathbf{x}_t) - v(0, \mathbf{x}_t)$$

- Remember that: $v(j, \mathbf{x}) = \pi(y, \mathbf{x}) + \delta \sum_{\mathbf{x}_{t+1}} V^\sigma(\mathbf{x}_{t+1}) f_x(\mathbf{x}_{t+1} | y, \mathbf{x})$.
Therefore:

$$\begin{aligned} \log P(j | \mathbf{x}_t) - \log P(k | \mathbf{x}_t) &= \pi(j, \mathbf{x}_t) - \pi(k, \mathbf{x}_t) + \\ &+ \delta \sum_{\mathbf{x}_{t+1}} [v(0, \mathbf{x}_{t+1}) - \log P(0 | \mathbf{x}_{t+1})] [f_x(\mathbf{x}_{t+1} | j, \mathbf{x}_t) - f_x(\mathbf{x}_{t+1} | k, \mathbf{x}_t)] \end{aligned}$$

FINITE DEPENDENCE REPRESENTATION

(3/4)

- Since $v(0, \mathbf{x}_{t+1}) = \pi(0, \mathbf{x}_{t+1}) + \delta \sum_{\mathbf{x}_{t+2}} V^\sigma(\mathbf{x}_{t+2}) f_x(\mathbf{x}_{t+2} | 0, \mathbf{x}_{t+1})$, we have that:

$$\begin{aligned} \log P(j | \mathbf{x}_t) - \log P(k | \mathbf{x}_t) &= \pi(j, \mathbf{x}_t) - \pi(k, \mathbf{x}_t) + \\ &+ \delta \sum_{\mathbf{x}_{t+1}} [\pi(0, \mathbf{x}_{t+1}) - \log P(0 | \mathbf{x}_{t+1})] [f_x(\mathbf{x}_{t+1} | j, \mathbf{x}_t) - f_x(\mathbf{x}_{t+1} | k, \mathbf{x}_t)] \\ &+ \delta^2 \sum_{\mathbf{x}_{t+1}} \sum_{\mathbf{x}_{t+2}} V^\sigma(\mathbf{x}_{t+2}) f_x(\mathbf{x}_{t+2} | 0, \mathbf{x}_{t+1}) [f_x(\mathbf{x}_{t+1} | j, \mathbf{x}_t) - f_x(\mathbf{x}_{t+1} | k, \mathbf{x}_t)] \end{aligned}$$

- There is a general class of dynamic models **[one-period finite dependence]** where this term is zero.

e.g., occupational choice; market entry-exit; machine replacement; dynamic demand of differentiated products; etc.

FINITE DEPENDENCE REPRESENTATION

(4/4)

- Under one-period finite dependence, the following condition holds at every-period t , any \mathbf{x}_t , and any pair (j, k) :

$$\begin{aligned} \log P(j|\mathbf{x}_t) - \log P(k|\mathbf{x}_t) &= \pi(j, \mathbf{x}_t; \boldsymbol{\theta}_\pi) - \pi(k, \mathbf{x}_t; \boldsymbol{\theta}_\pi) \\ &\quad - \delta \sum_{\mathbf{x}_{t+1}} \log P(0|\mathbf{x}_{t+1}) [f_x(\mathbf{x}_{t+1}|j, \mathbf{x}_t) - f_x(\mathbf{x}_{t+1}|k, \mathbf{x}_t)] \\ &\quad + \delta \sum_{\mathbf{x}_{t+1}} \pi(0, \mathbf{x}_{t+1}; \boldsymbol{\theta}_\pi) [f_x(\mathbf{x}_{t+1}|j, \mathbf{x}_t) - f(\mathbf{x}_{t+1}|k, \mathbf{x}_t)] \end{aligned}$$

- This equation provides moment conditions that can be used to estimate consistently the vector of parameters $\boldsymbol{\theta}_\pi$.
- Given a Nonparametric estimator of the reduced-form CCPs $P(j|\mathbf{x}_t)$, we can estimate structural parameters using a simple two-step GMM estimator.

EXAMPLE: MARKET ENTRY & EXIT

- $\mathbf{x}_t = (y_{t-1}, s_t)$; $\pi(0, \mathbf{x}_t) = 0$; $\pi(1, \mathbf{x}_t) = \theta_1 + \theta_2 s_t + \theta_3(1 - y_{t-1})$.
- This model has the one-period dependence property.
- We have:

$$\begin{aligned} \log P(1|\mathbf{x}_t) - \log P(0|\mathbf{x}_t) &= \theta_1 + \theta_2 s_t + \theta_3(1 - y_{t-1}) \\ -\delta \sum_{s_{t+1}} [\log P(0|1, s_{t+1}) - \log P(0|0, s_{t+1})] & f_s(s_{t+1}|s_t) \end{aligned}$$