

## TUT2: Production Function Estimation

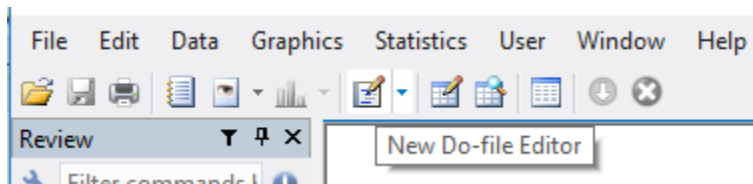
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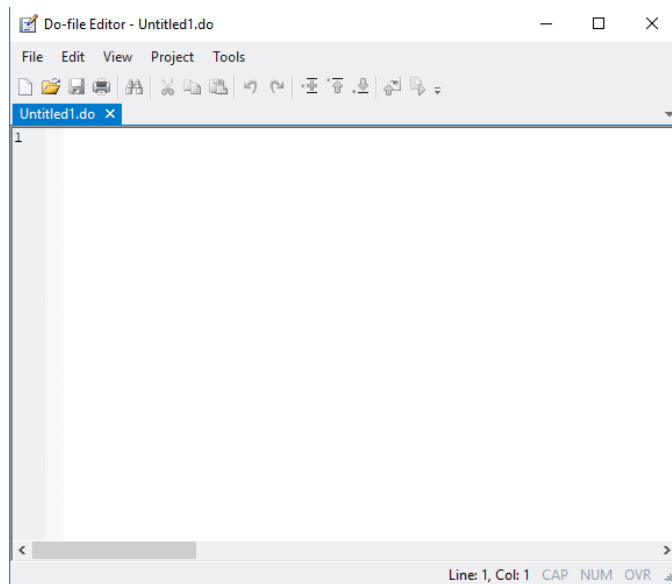
### 1. Create a do-file

The do-file contains the Stata commands that you wish to execute. Executing a do-file is the same as executing a series of commands interactively, only you have a permanent record of your commands. This allows you to quickly reproduce work you have already done and go from there.

Click the button shown as follows:



The do-file editor should open in a new window, with a clean page looking something like this:



### 2. Load the Blundell-Bond panel dataset

```
use "C:\Users\admin\Downloads\blundell_bond_2000_production_function.dta", clear
```

This directory depends on where you save the dataset. By typing "clear", it specifies that it is okay to replace the data in memory, even though the current data have not been saved to disk

### 3. Pooled OLS

#### 1) The simple version

**reg ln\_sales ln\_capital ln\_labor**

. reg ln\_sales ln\_capital ln\_labor

Source	SS	df	MS	Number of obs	=	4,072
Model	15942.9273	2	7971.46365	F(2, 4069)	=	63804.90
Residual	508.360451	4,069	.124934984	Prob > F	=	0.0000
				R-squared	=	0.9691
				Adj R-squared	=	0.9691
Total	16451.2878	4,071	4.04109255	Root MSE	=	.35346

ln_sales	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
ln_capital	.4298586	.0079525	54.05	0.000	.4142675 .4454498
ln_labor	.560581	.0096412	58.14	0.000	.541679 .5794829
_cons	3.005052	.0293099	102.53	0.000	2.947588 3.062515

#### 2) Adding time fixed effect into OLS regression

**reg ln\_sales ln\_capital ln\_labor i.year**

. reg ln\_sales ln\_capital ln\_labor i.year

Source	SS	df	MS	Number of obs	=	4,072
Model	15946.3907	9	1771.82119	F(9, 4062)	=	14254.66
Residual	504.897075	4,062	.124297655	Prob > F	=	0.0000
				R-squared	=	0.9693
				Adj R-squared	=	0.9692
Total	16451.2878	4,071	4.04109255	Root MSE	=	.35256

ln_sales	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
ln_capital	.4322828	.0081396	53.11	0.000	.4163247 .4482409
ln_labor	.5578836	.0098286	56.76	0.000	.5386142 .577153
year					
1983	-.0568626	.022107	-2.57	0.010	-.1002045 -.0135206
1984	-.050041	.0221342	-2.26	0.024	-.0934362 -.0066458
1985	-.0875714	.0221985	-3.94	0.000	-.1310926 -.0440503
1986	-.092866	.0222691	-4.17	0.000	-.1365256 -.0492063
1987	-.0580931	.0223043	-2.60	0.009	-.1018218 -.0143644
1988	-.0211632	.0223277	-0.95	0.343	-.0649378 .0226114
1989	-.0382923	.0224365	-1.71	0.088	-.0822802 .0056957
_cons	3.046843	.0315266	96.64	0.000	2.985033 3.108652

We would like to control for time effects whenever unexpected variation or special events may affect the outcome variable. However, we do not have to create dummy variable manually by using “gen” command. By typing “i.year” when using regressions, we are controlling time fixed-effects by creating dummy variables for each year from a categorical variable: year.

Note: in the above table, year 1982 is omitted. By default, the first (smallest) value will be used as reference category.

### 3) Use clustered standard error

**reg ln\_sales ln\_capital ln\_labor i.year, vce(cluster id)**

. reg ln\_sales ln\_capital ln\_labor i.year, vce(cluster id)

```
Linear regression               Number of obs   =    4,072
                              F(9, 508)       =   2507.63
                              Prob > F              =    0.0000
                              R-squared             =    0.9693
                              Root MSE         =    .35256
```

(Std. Err. adjusted for 509 clusters in id)

ln_sales	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
ln_capital	.4322828	.0274846	15.73	0.000	.3782853	.4862803
ln_labor	.5578836	.0308763	18.07	0.000	.4972227	.6185445
year						
1983	-.0568626	.0083657	-6.80	0.000	-.0732982	-.0404269
1984	-.050041	.0110933	-4.51	0.000	-.0718355	-.0282465
1985	-.0875714	.0135255	-6.47	0.000	-.1141442	-.0609987
1986	-.092866	.016461	-5.64	0.000	-.125206	-.0605259
1987	-.0580931	.0174944	-3.32	0.001	-.0924634	-.0237228
1988	-.0211632	.0185846	-1.14	0.255	-.0576754	.015349
1989	-.0382923	.020265	-1.89	0.059	-.0781058	.0015213
_cons	3.046843	.0915369	33.29	0.000	2.867005	3.22668

It is very unlikely that all observations in a data set are unrelated, but drawn from identical distributions. Some phenomena do not affect observations individually, but they affect groups of observations uniformly within each group. By using “vce(cluster id)”, we allow for correlation between observations. Clustered standard error will increase your confidence intervals. The higher the clustering level, the larger the resulting standard error. Hence, less stars in your tables.

### 4) Add lagged values as explanatory variables

**reg ln\_sales ln\_capital ln\_labor l.ln\_sales l.ln\_capital l.ln\_labor i.year, vce(cluster id)**

```

Linear regression              Number of obs   =    3,563
                              F(11, 508)     =   75113.93
                              Prob > F           =    0.0000
                              R-squared         =    0.9949
                              Root MSE      =    .1426

```

(Std. Err. adjusted for 509 clusters in id)

ln_sales	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
ln_capital	.2348098	.0352859	6.65	0.000	.1654854	.3041341
ln_labor	.4789653	.0288963	16.58	0.000	.4221944	.5357362
ln_sales L1.	.9216418	.0105283	87.54	0.000	.9009575	.9423261
ln_capital L1.	-.2120621	.0347141	-6.11	0.000	-.280263	-.1438613
ln_labor L1.	-.4233997	.0305806	-13.85	0.000	-.4834797	-.3633198
year						
1984	.0661134	.0096619	6.84	0.000	.0471313	.0850956
1985	.0127151	.0101712	1.25	0.212	-.0072678	.032698
1986	.034839	.0102827	3.39	0.001	.0146373	.0550408
1987	.0767564	.0086928	8.83	0.000	.0596782	.0938346
1988	.0779253	.0094209	8.27	0.000	.0594167	.096434
1989	.0356445	.0097659	3.65	0.000	.0164579	.0548311
_cons	.246604	.0322971	7.64	0.000	.1831516	.3100564

By using “l.variables”, we are creating lagged values of those variables since expect the effects of labour/capital on outputs to appear with a delay. That is, this year's value of output may depend on last year's value of labour/capital/output rather than on the current value.

#### 4. Within Groups (or fixed effects estimator)

A variety of commands are available for estimating fixed effects regressions. The most efficient method is the fixed effects regression (within estimation).

Firstly, we want to declare data to be panel data; panel variable entered at first, time variable follows:

***xtset id year***

```

xtset id year
  panel variable:  id (strongly balanced)
  time variable:  year, 1982 to 1989
  delta: 1 unit

```

Notes: The terms balanced and unbalanced are often used to describe whether a panel dataset is missing some observations. If a dataset does not contain a time variable, then panels are considered balanced if each panel contains the same number of observations; otherwise, the panels are unbalanced.

When the dataset contains a time variable, panels are said to be strongly balanced if each panel contains the same time points, weakly balanced if each panel contains the same number of observations but not the same time points, and unbalanced otherwise.

Stata's `xtreg` command is built for panel data regressions. Use the `fe` option to specify fixed effects:

**`xtreg ln_sales ln_labor ln_capital l.ln_sales l.ln_capital l.ln_labor i.year, fe vce(cluster id)`**

```
Fixed-effects (within) regression      Number of obs   =    3,563
Group variable: id                   Number of groups =    509

R-sq:                                Obs per group:
    within = 0.7825                    min =          7
    between = 0.9879                   avg =         7.0
    overall = 0.9847                   max =          7

                                F(11,508)        =    345.86
corr(u_i, Xb) = 0.7191                Prob > F        =    0.0000

                                (Std. Err. adjusted for 509 clusters in id)
```

ln_sales	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
ln_labor	.4880013	.0299207	16.31	0.000	.4292178	.5467848
ln_capital	.1765454	.0340853	5.18	0.000	.1095799	.2435109
ln_sales						
L1.	.4039344	.0293221	13.78	0.000	.3463269	.4615418
ln_capital						
L1.	-.1305487	.0252882	-5.16	0.000	-.1802311	-.0808663
ln_labor						
L1.	-.0231194	.03455	-0.67	0.504	-.0909977	.044759
year						
1984	.0564054	.0074048	7.62	0.000	.0418576	.0709533
1985	.0271379	.0092935	2.92	0.004	.0088795	.0453963
1986	.0494812	.0116301	4.25	0.000	.0266322	.0723302
1987	.1033078	.0116548	8.86	0.000	.0804104	.1262053
1988	.1310847	.01305	10.04	0.000	.1054461	.1567234
1989	.1174383	.0150838	7.79	0.000	.087804	.1470726
_cons	2.625541	.1591593	16.50	0.000	2.31285	2.938233
sigma_u	.31731619					
sigma_e	.12076713					
rho	.87347826	(fraction of variance due to u_i)				

In the above table, the interpretations of all Coef. Standard error, T statistics, P value, Confidence interval are same as OLS regression's output. At the bottom, 'rho' represents how much of the variance is due to differences across panels. 'rho' is known as the intraclass correlation.

## 5. First differences GMM (or Arellano-Bond estimator)

### 1) Search and install the command

Two Arellano-Bond estimators are available for Stata 9.0 – one incorporated into Stata 9+ (called `xtabond`) and one propriator program written by Roodman (2006) (called `xtabond2`). `xtabond2` can do everything that `xtabond` does and has many additional features. Since `xtabond2` is not an official command of Stata 9+, it has to be downloaded. By searching the command:

#### ***search xtabond2***

```
search for xtabond2

Search of official help files, FAQs, Examples, SJs, and STBs

SJ-12-4 st0159_1 . . . . . Software update for xtabond2
(help xtabond2 if installed) . . . . . D. Roodman
Q4/12 SJ 12(4):766--767
bug fixes and added features such as support for factor
variables

SJ-9-1 st0159 . . How to do xtabond2: An intro. to difference and system GMM
(help xtabond2 if installed) . . . . . D. Roodman
Q1/09 SJ 9(1):86--136
introduces linear GMM; describes how limited time span
and potential for fixed effects and endogenous regressors
drive the design of estimators; shows how to apply
difference and system GMM estimators with xtabond2
```

Then by clicking the package and click “click here to install”, you will complete the installment of the command:

```
package st0159_1 from http://www.stata-journal.com/software/sj12-4

TITLE
    SJ12-4 st0159_1. Update: 'XTABOND2': module to...

DESCRIPTION/AUTHOR(S)
    Update: 'XTABOND2': module to extend xtabond
    dynamic panel data estimator
    by David Roodman, Center for Global Development
    Support: DRoodman@CGDEV.ORG
    After installation, type help xtabond2 and
    abar

INSTALLATION FILES (click here to install)
    st0159_1/xtabond2.ado
    st0159_1/xtabond2.mata
    st0159_1/xtabond2.hlp
    st0159_1/xtab2_p.ado
    st0159_1/xtabond2.mlib
    st0159_1/abar.hlp
    st0159_1/abar.ado
```

Once you install “`xtabond2`”, the following command shows the help file if you are not sure how to use it:

#### ***help xtabond2***

## 2) Use Lag 2 + as instruments

The command `xtabond2` is followed by the dependent variable (`inv`) and the list of all right-hand-side variables:

```
xtabond2 ln_sales ln_capital ln_labor l.ln_sales l.ln_capital l.ln_labor i.year, gmm(ln_sales ln_capital ln_labor, lag(2.)) iv(i.year) robust nolevelq
```

After the comma are given two lists of variables. `gmm( )` (or `gmmstyle( )`) lists the endogenous variables, which are instrumented with GMM-style instruments, i.e. lagged values of the variables in levels: Y,K,L. With `lag(2. )` we have instructed Stata to use the second lag and following lags of the endogenous variables as instruments. The second lag is required, because it is not correlated with the current error term.

The second list of explanatory variables, `iv( )` (or `ivstyle( )`), lists all strictly exogenous variables as well as the additional instrumental variables, therefore, are not listed before the comma in the Stata command. What this option essentially does for the included exogenous variables is tell Stata to use the variables themselves as their own instruments.

`Robust` specifies that the resulting standard errors are consistent with panel-specific autocorrelation and heteroskedasticity in one-step estimation. `Nolevel` (or `noleveleq`) tells Stata to apply the difference GMM estimator. By default, `xtabond2` will apply the system GMM, if you don't specify `nolevel`. (System GMM is discussed in section 6)

```
Dynamic panel-data estimation, one-step difference GMM
```

Group variable: id		Number of obs = 3054				
Time variable: year		Number of groups = 509				
Number of instruments = 69		Obs per group: min = 6				
Wald chi2(13) = 1340.23		avg = 6.00				
Prob > chi2 = 0.000		max = 6				
ln_sales	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
ln_capital	.1318291	.1179029	1.12	0.264	-.0992562	.3629145
ln_labor	.5128697	.0892432	5.75	0.000	.3379563	.6877831
ln_sales L1.	.3264209	.0521606	6.26	0.000	.2241881	.4286538
ln_capital L1.	-.2066298	.0949589	-2.18	0.030	-.3927458	-.0205137
ln_labor L1.	.0726061	.0927269	0.78	0.434	-.1091353	.2543475
year	0 (empty)					
1982	-.1899966	.0285349	-6.66	0.000	-.245924	-.1340691
1983	-.122387	.0231032	-5.30	0.000	-.1676683	-.0771056
1984	-.1334924	.017638	-7.57	0.000	-.1680623	-.0989225
1985	-.0986976	.0134461	-7.34	0.000	-.1250514	-.0723437
1986	-.0347582	.0108629	-3.20	0.001	-.0560491	-.0134673
1987	.0028642	.0093652	0.31	0.760	-.0154913	.0212196
1988	0 (omitted)					
1989	0 (omitted)					

```
Instruments for first differences equation
Standard
D.(1982b.year 1983.year 1984.year 1985.year 1986.year 1987.year 1988.year 1989.year)
GMM-type (missing=0, separate instruments for each period unless collapsed)
L(2/7).(ln_sales ln_capital ln_labor)

Arellano-Bond test for AR(1) in first differences: z = -6.21 Pr > z = 0.000
Arellano-Bond test for AR(2) in first differences: z = -1.36 Pr > z = 0.173

Sargan test of overid. restrictions: chi2(56) = 213.56 Prob > chi2 = 0.000
(Not robust, but not weakened by many instruments.)
Hansen test of overid. restrictions: chi2(56) = 98.39 Prob > chi2 = 0.000
(Robust, but weakened by many instruments.)

Difference-in-Hansen tests of exogeneity of instrument subsets:
iv(1982b.year 1983.year 1984.year 1985.year 1986.year 1987.year 1988.year 1989.year)
Hansen test excluding group: chi2(50) = 78.58 Prob > chi2 = 0.006
Difference (null H = exogenous): chi2(6) = 19.81 Prob > chi2 = 0.003
```

By default Stata reports three additional tests: Sargan test, AR(1) and AR(2) tests. The Sargan test has a null hypothesis of “the instruments as a group are exogenous”. Therefore, the higher the p-value of the Sargan statistic the better. In robust estimation Stata reports the Hansen J statistic instead of the Sargan with the same null hypothesis.

The Arellano – Bond test for autocorrelation has a null hypothesis of no autocorrelation and is applied to the differenced residuals. The test for AR (2) in first differences is more important, because it will detect autocorrelation in levels.

### 3) Use Lag 3+ as instruments

The following command `lag(3. )` omits the levels of the variables dated t-2 from the set of instruments:

**`xtabond2 ln_sales ln_capital ln_labor l.ln_sales l.ln_capital l.ln_labor i.year, gmm(ln_sales ln_capital ln_labor, lag(3.)) iv(i.year) robust nolevelq`**

```
Group variable: id                Number of obs   =   3054
Time variable : year            Number of groups =   509
Number of instruments = 51      Obs per group:  min =    6
Wald chi2(13) = 1665.43        avg =   6.00
Prob > chi2 = 0.000           max =    6
```

	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
ln_sales						
ln_capital	.1940721	.153916	1.26	0.207	-.1075977	.4957419
ln_labor	.4987448	.101461	4.92	0.000	.2998848	.6976047
ln_sales						
L1.	.4261318	.0791759	5.38	0.000	.27095	.5813137
ln_capital						
L1.	-.1054642	.109791	-0.96	0.337	-.3206507	.1097223
ln_labor						
L1.	-.1469956	.113259	-1.30	0.194	-.3689792	.074988
year						
1982	0	(empty)				
1983	-.092768	.0294955	-3.15	0.002	-.150578	-.0349579
1984	-.0374932	.0211068	-1.78	0.076	-.0788618	.0038754
1985	-.0676855	.0144321	-4.69	0.000	-.0959719	-.0393991
1986	-.0491738	.0101795	-4.83	0.000	-.0691252	-.0292224
1987	0	(omitted)				
1988	.0268377	.0081829	3.28	0.001	.0107994	.0428759
1989	.0115639	.0149883	0.77	0.440	-.0178126	.0409405

```
Instruments for first differences equation
Standard
D.(1982b.year 1983.year 1984.year 1985.year 1986.year 1987.year 1988.year
1989.year)
GMM-type (missing=0, separate instruments for each period unless collapsed)
L(3/7).(ln_sales ln_capital ln_labor)
```

```
Arellano-Bond test for AR(1) in first differences: z = -5.09 Pr > z = 0.000
Arellano-Bond test for AR(2) in first differences: z = -0.79 Pr > z = 0.429
```

```
Sargan test of overid. restrictions: chi2(38) = 86.43 Prob > chi2 = 0.000
(Not robust, but not weakened by many instruments.)
```

```
Hansen test of overid. restrictions: chi2(38) = 53.66 Prob > chi2 = 0.047
(Robust, but weakened by many instruments.)
```

```
Difference-in-Hansen tests of exogeneity of instrument subsets:
iv(1982b.year 1983.year 1984.year 1985.year 1986.year 1987.year 1988.year 1989.year)
Hansen test excluding group: chi2(32) = 45.74 Prob > chi2 = 0.055
Difference (null H = exogenous): chi2(6) = 7.92 Prob > chi2 = 0.244
```



## 6. System GMM (or Blundell-Bond estimator)

Sometimes the lagged levels of the regressors are poor instruments for the first-differenced regressors. In this case, one should use the augmented version – “system GMM”. The system GMM estimator uses the levels equation to obtain a system of two equations: one differenced and one in levels. By adding the second equation additional instruments can be obtained. Therefore, the variables in levels in the second equation are instrumented with their own first differences. The command is following:

```
xtabond2 ln_sales ln_capital ln_labor l.ln_sales l.ln_capital l.ln_labor i.year, gmm(ln_sales ln_capital ln_labor, lag(2 .)) iv(i.year, equation(level)) robust h(1)
```

nolevel is not included after the comma in the command and Stata defaults to the system GMM. The h(1) option uses 2SLS as the one-step estimator, which is the value in the original implementation of the system GMM estimator in Blundell and Bond(1998).

The gmm(, lag(2 .)) option uses the lagged levels of Y, L and K dated t-2 and earlier as instruments for the equations in first-differences; and (correspondingly) the lagged first-differences of Y, L and K dated t-1 (only) as instruments for the equations in levels. This is the default specification of gmm-style instruments for the levels equations.

‘xtabond2’ offers the equation () sub-option, which specifies which equation should use the instruments: first-difference only (equation (diff)) or levels only (equation (level)). The default is both equations. In this case, the iv(i.year, equation(level)) option uses the year dummies as instruments for the equations in levels only.

Alternatively, using the following code, we choose lagged level from t-3 as instruments:

```
xtabond2 ln_sales ln_capital ln_labor l.ln_sales l.ln_capital l.ln_labor i.year, gmm(ln_sales ln_capital ln_labor, lag(3 .)) iv(i.year, equation(level)) robust h(1)
```

The gmm(, lag(3 .)) option uses the lagged levels of Y, L and K dated t-3 and earlier as instruments for the equations in first-differences; and (correspondingly) the lagged first-differences of Y, L and K dated t-2 (only) as instruments for the equations in levels.

The following tables are generated by the first type of code, i.e. lag(2):

```

Group variable: id                Number of obs   =    3563
Time variable : year            Number of groups =    509
Number of instruments = 88      Obs per group: min =     7
Wald chi2(13) = 404276.63      avg           =    7.00
Prob > chi2   =     0.000      max           =     7

```

ln_sales	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
ln_capital	.3666974	.130163	2.82	0.005	.1115826	.6218122
ln_labor	.623971	.1078013	5.79	0.000	.4126843	.8352578
ln_sales						
L1.	.4650298	.0520106	8.94	0.000	.3630909	.5669688
ln_capital						
L1.	-.3304688	.1063253	-3.11	0.002	-.5388625	-.1220752
ln_labor						
L1.	-.0925121	.109246	-0.85	0.397	-.3066303	.121606
year						
1982	0	(empty)				
1983	2.218975	.2356304	9.42	0.000	1.757147	2.680802
1984	2.266482	.2368093	9.57	0.000	1.802344	2.730619
1985	2.242476	.2399846	9.34	0.000	1.772115	2.712837
1986	2.271655	.2424574	9.37	0.000	1.796447	2.746863
1987	2.32348	.2427699	9.57	0.000	1.84766	2.799301
1988	2.348755	.2448667	9.59	0.000	1.868825	2.828685
1989	2.322886	.249252	9.32	0.000	1.834361	2.811411
_cons	0	(omitted)				

```

Instruments for first differences equation
GMM-type (missing=0, separate instruments for each period unless collapsed)
L(2/7).(ln_sales ln_capital ln_labor)
Instruments for levels equation
Standard
1982b.year 1983.year 1984.year 1985.year 1986.year 1987.year 1988.year
1989.year
_cons
GMM-type (missing=0, separate instruments for each period unless collapsed)
DL.(ln_sales ln_capital ln_labor)

```

```

Arellano-Bond test for AR(1) in first differences: z = -8.17 Pr > z = 0.000
Arellano-Bond test for AR(2) in first differences: z = -0.60 Pr > z = 0.547

```

```

Sargan test of overid. restrictions: chi2(74) = 118.62 Prob > chi2 = 0.001
(Not robust, but not weakened by many instruments.)
Hansen test of overid. restrictions: chi2(74) = 142.10 Prob > chi2 = 0.000
(Robust, but weakened by many instruments.)

```

```

Difference-in-Hansen tests of exogeneity of instrument subsets:
GMM instruments for levels
Hansen test excluding group: chi2(56) = 97.03 Prob > chi2 = 0.001
Difference (null H = exogenous): chi2(18) = 45.07 Prob > chi2 = 0.000
iv(1982b.year 1983.year 1984.year 1985.year 1986.year 1987.year 1988.year 1989.year, eq(level))
Hansen test excluding group: chi2(68) = 122.25 Prob > chi2 = 0.000
Difference (null H = exogenous): chi2(6) = 19.85 Prob > chi2 = 0.003

```

The following tables are generated by the first type of code, i.e. lag(3):

